



### MACROECONOMETRIC MODELS II

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# Q&A

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# SINGLE-EQUATION ECONOMETRICS—A REMINDER

Summary

## A SIMPLE EXAMPLE: NATIONAL ECONOMY

Definition Y: Gross Domestic Product	Y(t) = C(t) + I(t) + G(t) + X(t) - M(t)
Consumption function of private households C	$C(t) = \alpha 0 + \alpha 1 * Y(t-1)$
Private Investment function I	$I(t) = \lambda * (Y(t) - Y(t-1))$
Government spending G	G (t): exogenous
Exports X	X(t): exogenous
Imports M	$M(t) = \gamma 0 + y 1 * Y(t)$

# SUMMARY OF SINGLE EQUATION

Theoretical Model	$Y = \beta 1 + \beta 2 * X + u$
Statistical model: Fitted values	$\hat{Y} = b_1 + b_2 X$
OLS Estimator of b1 (Intercept)	$\boldsymbol{b}_1 = \overline{Y} - \boldsymbol{b}_2 \overline{X}$
OLS Estimator of b2 (slope)	$b_2 = \frac{\sum (X_i - \overline{X})(Y_i - \overline{Y})}{\sum (X_i - \overline{X})^2}$
Goodness of fit:	$R^{2} = \frac{ESS}{TSS} = \frac{\sum (\hat{Y}_{i} - \overline{Y})^{2}}{\sum (Y_{i} - \overline{Y})^{2}}$

# MODEL SOFTWARE: gretl

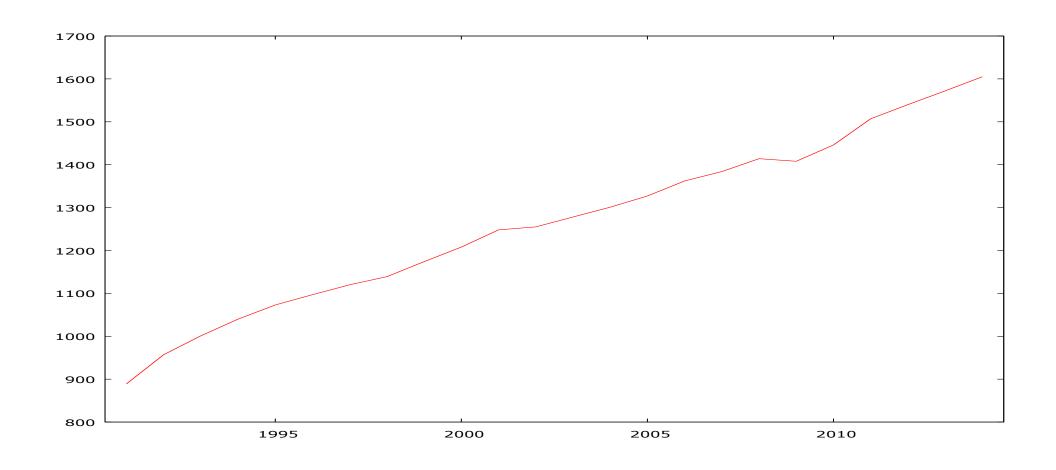
• gretl: Gnu Regression, Econometrics and Time-series Library

- Source in Internet for a free download
- http://gretl.sourceforge.net/

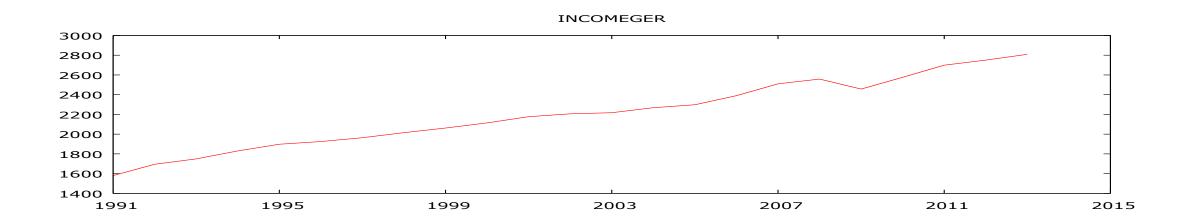
# SINGLE EQUATION ESTIMATION

- Application of "gretl" econometric software
- Consumption function
- Investment function
- Import function

# Consumption, Germany Time Series



# GDP, Germany, Time Series



# **Consumption Function**

- Modell 1: KQ, benutze die Beobachtungen 1992-2014 (T = 23)
- Abhängige Variable: CONSUMGER

```
Koeffizient Std. Fehler t-Quotient p-Wert
                                                     ***
          100,847
                     28,4018
                               3,5507
                                          0.0019

    const

                    0,534741 0,0127219 42,0333
                                                                ***

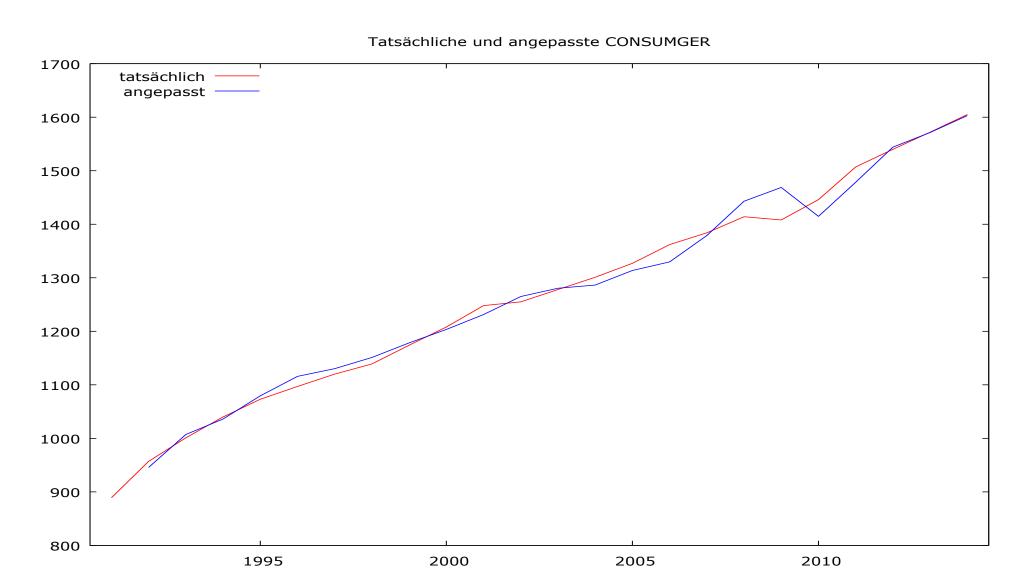
    INCOMEGER 1

                                                     <0,0001
  Mittel d. abh. Var.
                     1280,696
                                          Stdabw. d. abh. Var.
                                                                187,2959
• Summe d. quad. Res.
                                9065,268
                                                     Stdfehler d. Regress.
                                                                           20,77690
  R-Quadrat
                     0,988254
                                          Korrigiertes R-Quadrat 0,987694
• F(1, 21) 1766,796
                                          9,37e-22
                                P-Wert(F)

    Log-Likelihood

                                          Akaike-Kriterium
                     -101,3678
                                                                206.7355
• Schwarz-Kriterium 209,0065
                                          Hannan-Quinn-Kriterium
                                                                           207,3067
• rho
           0,169094
                               Durbin-Watson-Stat
                                                     1,647511
```

# Goodness of fit



# MULTI-EQUATION ECONOMETRICS: METHODS AND PROBLEMS

## MODEL ESTIMATION

Application of "gretl" econometric software

- BEHAVIOURAL EQUATIONS
- Consumption function
- Investment function
- Import function
- Government consumption

- EXOGENOUS VARIABLES
- Government consumption
- Exports
- DEFINITION
- Gross domestic product

# Model Results I

```
cret1 Output fills MS 2015-09-01 03:39. Seite 1
Gleichungstysten. Scheinbar unverbundene Regressionen
Gleichung 1: SUR, bemutze die Beobachtungen 1992-2013 (T = 22)
Abbinoice Variable: CONSUMGER
             Koeffizient Std.-fehler t-Quotient p-Wert
  INCOMMECSER 0.564301 0.00155185 363.6 2.21-041 ***
Mittel d. abb. Var. 1265,955 Stdabw. d. abb. Var. 177,5228
Summe d. cound. Res. 6605.304 Stdfebler d. Recress. 17.32747
Gleichung 2: SUR, bemutze die Beobachtungen 1992-2013 (T = 22)
Mohimorica Variable: DIVESTGER
```

# Model Results II

```
Koeffizient Std.-fehler t-Quotient p-Wert
 d_INCOMEGER 5,33774 0,830885 6,424 2,29e-06 ***
Mittel d. abb. Var. 471,6364 Stdabw. d. abb. Var. 42,33151
Summe d. quad. Res. 1892414 Stdfebler d. Regress. 293,2897
Gleichung 3: SUR, benutze die Beobachtungen 1992-2013 (T = 22)
Abhinorice Variable: DAFORTGER
            Koeffizient Std.-fehler t-Quotient p-Wert
 INCOMEGER 0.323789 0.0136547 23.71 1.21--016 ***
Mittel d. abb. Var. 694,0455 Stdabw. d. abb. Var. 250,0664
Summe d. quad. Res. 464640,7 Stdfebler d. Regress. 145,3273
```

# **Model Results III**

```
Gleichungzübergreifende Regidnen-Var-Kov-Mat
(Über der Diagonalen Korrelationen)

300,24 (0,492) (-0,558)
2498,2 86019, (-0,151)
-1404,5 -6422,5 21120,

log-Determinante = 26,346
Breugch-Pagan-Test für diagonale Kovariansmatrix:
Chi-Quadrat(3) = 12,6597 [0,0054]
```

Cf: Annex:

Presentation based on

**Dougherty: Introduction to Econometrics 4e** 

http://global.oup.com/uk/orc/busecon/economics/dougherty4e/

# Q&A

# ANNEX

Presentation based on

**Dougherty: Introduction to Econometrics 4e** 

http://global.oup.com/uk/orc/busecon/economics/dougherty4e/

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

This sequence shows why OLS is likely to yield inconsistent estimates in models composed of two or more simultaneous relationships.

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

In this example we suppose that we have data on p, the annual rate of price inflation, w, the annual rate of wage inflation, and U, the rate of unemployment, for a sample of countries.

$$p = \beta_1 + \beta_2 w + u_p$$

$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

We hypothesize that increases in wages lead to increases in prices and so p is positively influenced by  $w(\beta_2 > 0)$ .

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

We also suppose that workers try to protect their real wages by negotiating for increases in wages as prices rise, but their ability to so is the weaker, the greater is the rate of unemployment ( $\alpha_2 > 0$ ,  $\alpha_3 < 0$ ).

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

endogenous: p, w

The model involves some circularity, in that w is a determinant of p, and p is a determinant of w. Variables whose values are determined interactively within the model are described as endogenous variables.

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

endogenous: p, w

exogenous:

We will cut through the circularity by expressing p and w in terms of their ultimate determinants, U and the disturbance terms  $u_p$  and  $u_{w^*}$  Variables such as U whose values are determined outside the model are described as exogenous variables.

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = \beta_1 + \beta_2 (\alpha_1 + \alpha_2 p + \alpha_3 U + u_w) + u_p$$

We will start with p. The first step is to substitute for w from the second equation.

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = \beta_1 + \beta_2 (\alpha_1 + \alpha_2 p + \alpha_3 U + u_w) + u_p$$
$$(1 - \alpha_2 \beta_2) p = \beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w$$

$$p = \frac{\beta_{1} + \alpha_{1}\beta_{2} + \alpha_{3}\beta_{2}U + u_{p} + \beta_{2}u_{w}}{1 - \alpha_{2}\beta_{2}}$$

We bring the terms involving p together on the left side of the equation and thus express p in terms of U,  $u_p$ , and  $u_w$ .

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = \beta_1 + \beta_2 (\alpha_1 + \alpha_2 p + \alpha_3 U + u_w) + u_p$$
$$(1 - \alpha_2 \beta_2) p = \beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w$$

$$p = \frac{\beta_{1} + \alpha_{1}\beta_{2} + \alpha_{3}\beta_{2}U + u_{p} + \beta_{2}u_{w}}{1 - \alpha_{2}\beta_{2}}$$

$$w = \alpha_1 + \alpha_2 \left( \beta_1 + \beta_2 w + u_p \right) + \alpha_3 U + u_w$$

Next we take the equation for w and substitute for p from the first equation.

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = \beta_1 + \beta_2 (\alpha_1 + \alpha_2 p + \alpha_3 U + u_w) + u_p$$
$$(1 - \alpha_2 \beta_2) p = \beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w$$

$$p = \frac{\beta_{1} + \alpha_{1}\beta_{2} + \alpha_{3}\beta_{2}U + u_{p} + \beta_{2}u_{w}}{1 - \alpha_{2}\beta_{2}}$$

$$w = \alpha_1 + \alpha_2 \left( \beta_1 + \beta_2 w + u_p \right) + \alpha_3 U + u_w$$
$$(1 - \alpha_2 \beta_2) w = \alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w$$

$$w = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}$$

We bring the terms involving w together on the left side of the equation and thus express w in terms of U,  $u_p$ , and  $u_w$ .

$$p = \beta_1 + \beta_2 w + u_p$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = \beta_1 + \beta_2 (\alpha_1 + \alpha_2 p + \alpha_3 U + u_w) + u_p$$
$$(1 - \alpha_2 \beta_2) p = \beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w$$

$$p = \frac{\beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w}{1 - \alpha_2 \beta_2}$$

$$w = \alpha_1 + \alpha_2 (\beta_1 + \beta_2 w + u_p) + \alpha_3 U + u_w$$
$$(1 - \alpha_2 \beta_2) w = \alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w$$

$$w = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}$$

The original equations, representing the economic relationships among the variables, are described as the structural equations.

$$p = \beta_1 + \beta_2 w + u_1$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = \beta_1 + \beta_2 (\alpha_1 + \alpha_2 p + \alpha_3 U + u_w) + u_p$$
$$(1 - \alpha_2 \beta_2) p = \beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w$$

reduced form equation

$$p = \frac{\beta_{1} + \alpha_{1}\beta_{2} + \alpha_{3}\beta_{2}U + u_{p} + \beta_{2}u_{w}}{1 - \alpha_{2}\beta_{2}}$$

$$w = \alpha_1 + \alpha_2 (\beta_1 + \beta_2 w + u_p) + \alpha_3 U + u_w$$
$$(1 - \alpha_2 \beta_2) w = \alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w$$

reduced form equation

$$w = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}$$

The equations expressing the endogenous variables in terms of the exogenous variable(s) and the disturbance terms are described as the reduced form equations.

$$p = \beta_1 + \beta_2 w + u_1$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = \beta_1 + \beta_2 (\alpha_1 + \alpha_2 p + \alpha_3 U + u_w) + u_p$$
$$(1 - \alpha_2 \beta_2) p = \beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w$$

reduced form equation

$$p = \frac{\beta_{1} + \alpha_{1}\beta_{2} + \alpha_{3}\beta_{2}U + u_{p} + \beta_{2}u_{w}}{1 - \alpha_{2}\beta_{2}}$$

$$w = \alpha_1 + \alpha_2 (\beta_1 + \beta_2 w + u_p) + \alpha_3 U + u_w$$
$$(1 - \alpha_2 \beta_2) w = \alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w$$

reduced form equation

$$w = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}$$

The reduced form equations have two important roles. They can indicate that we have a serious econometric problem, but they may also provide a solution to it.

$$p = \beta_1 + \beta_2 w + \frac{u_p}{u_p}$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = \beta_1 + \beta_2 (\alpha_1 + \alpha_2 p + \alpha_3 U + u_w) + u_p$$
$$(1 - \alpha_2 \beta_2) p = \beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w$$

$$p = \frac{\beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w}{1 - \alpha_2 \beta_2}$$

$$w = \alpha_1 + \alpha_2 \left( \beta_1 + \beta_2 w + u_p \right) + \alpha_3 U + u_w$$
$$(1 - \alpha_2 \beta_2) w = \alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w$$

$$\mathbf{w} = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 \mathbf{u}_p + \mathbf{u}_w}{1 - \alpha_2 \beta_2}$$

The problem is the violation of Assumption B.7 that the disturbance term be distributed independently of the explanatory variable(s). In the first equation, w has a component  $u_n$ . OLS would therefore yield inconsistent estimates if used to fit the equation.

$$p = \beta_1 + \beta_2 w + u$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = \beta_1 + \beta_2 (\alpha_1 + \alpha_2 p + \alpha_3 U + u_w) + u_p$$
$$(1 - \alpha_2 \beta_2) p = \beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w$$

$$p = \frac{\beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w}{1 - \alpha_2 \beta_2}$$

$$w = \alpha_1 + \alpha_2 (\beta_1 + \beta_2 w + u_p) + \alpha_3 U + u_w$$
$$(1 - \alpha_2 \beta_2) w = \alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w$$

$$w = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}$$

Likewise, in the second equation, p has a component  $u_{w}$ 

structural equations

$$p = \beta_1 + \beta_2 w + u_p$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$b_2^{\text{OLS}} = \frac{\sum (w_i - \overline{w})(p_i - \overline{p})}{\sum (w_i - \overline{w})^2}$$

We will investigate the sign of the bias in the slope coefficient if OLS is used to fit the price inflation equation.

structural equations

$$p = \beta_1 + \beta_2 w + u_p$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$b_2^{\text{OLS}} = \frac{\sum (w_i - \overline{w})(p_i - \overline{p})}{\sum (w_i - \overline{w})^2}$$

reduced form equation

$$p = \frac{\beta_{1} + \alpha_{1}\beta_{2} + \alpha_{3}\beta_{2}U + u_{p} + \beta_{2}u_{w}}{1 - \alpha_{2}\beta_{2}}$$

As usual we start by substituting for the dependent variable using the true model. For this purpose, we could use either the structural equation or the reduced form equation for p.

$$p = \beta_1 + \beta_2 w + u_p$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$b_2^{\text{OLS}} = \frac{\sum (w_i - \overline{w})(p_i - \overline{p})}{\sum (w_i - \overline{w})^2}$$

$$= \frac{\sum (w_i - \overline{w})([\beta_1 + \beta_2 w_i + u_{pi}] - [\beta_1 + \beta_2 \overline{w} + \overline{u}_p])}{\sum (w_i - \overline{w})^2}$$

reduced form equation

$$p = \frac{\beta_{1} + \alpha_{1}\beta_{2} + \alpha_{3}\beta_{2}U + u_{p} + \beta_{2}u_{w}}{1 - \alpha_{2}\beta_{2}}$$

The algebra is simpler if we use the structural equation.

structural equations

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$b_{2}^{\text{OLS}} = \frac{\sum (w_{i} - \overline{w})(p_{i} - \overline{p})}{\sum (w_{i} - \overline{w})^{2}}$$

$$= \frac{\sum (w_{i} - \overline{w})([\beta_{1} + \beta_{2}w_{i} + u_{pi}] - [\beta_{1} + \beta_{2}\overline{w} + \overline{u}_{p}])}{\sum (w_{i} - \overline{w})^{2}}$$

$$= \frac{\sum (w_{i} - \overline{w})(\beta_{2}[w_{i} - \overline{w}] + u_{pi} - \overline{u}_{p})}{\sum (w_{i} - \overline{w})^{2}}$$

The  $\beta_1$  terms cancel. We rearrange the rest of the second factor in the numerator.

structural equations

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$b_{2}^{\text{OLS}} = \frac{\sum (w_{i} - \overline{w})(p_{i} - \overline{p})}{\sum (w_{i} - \overline{w})^{2}}$$

$$= \frac{\sum (w_{i} - \overline{w})([\beta_{1} + \beta_{2}w_{i} + u_{pi}] - [\beta_{1} + \beta_{2}\overline{w} + \overline{u}_{p}])}{\sum (w_{i} - \overline{w})^{2}}$$

$$= \frac{\sum (w_{i} - \overline{w})(\beta_{2}[w_{i} - \overline{w}] + u_{pi} - \overline{u}_{p})}{\sum (w_{i} - \overline{w})^{2}}$$

$$= \beta_{2} + \frac{\sum (w_{i} - \overline{w})(u_{pi} - \overline{u}_{p})}{\sum (w_{i} - \overline{w})^{2}}$$

Hence we obtain the usual decomposition into true value and error term.

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$b_2^{\text{OLS}} = \beta_2 + \frac{\sum (w_i - \overline{w})(u_{pi} - \overline{u}_p)}{\sum (w_i - \overline{w})^2}$$

We will now investigate the properties of the error term. Of course, we would like it to have expected value 0, making the estimator unbiased.

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$b_2^{\text{OLS}} = \beta_2 + \frac{\sum (w_i - \overline{w})(u_{pi} - \overline{u}_p)}{\sum (w_i - \overline{w})^2}$$

$$\mathbf{w} = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 \mathbf{u}_p + \mathbf{u}_w}{1 - \alpha_2 \beta_2}$$

However, the error term is a nonlinear function of both  $u_p$  and  $u_w$  because both are components of w. As a consequence, it is not possible to obtain a closed-form analytical expression for its expected value.

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$b_2^{\text{OLS}} = \beta_2 + \frac{\sum (w_i - \overline{w})(u_{pi} - \overline{u}_p)}{\sum (w_i - \overline{w})^2}$$

$$\mathbf{w} = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 \mathbf{u}_p + \mathbf{u}_w}{1 - \alpha_2 \beta_2}$$

We will investigate the large-sample properties instead. We will demonstrate that the estimator is inconsistent, and this will imply that it has undesirable finite-sample properties.

$$p = \beta_1 + \beta_2 w + u_{\mu}$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$b_2^{\text{OLS}} = \beta_2 + \frac{\sum (w_i - \overline{w})(u_{pi} - \overline{u}_p)}{\sum (w_i - \overline{w})^2}$$

$$\mathbf{plim}\left(\frac{\sum (w_i - \overline{w})(u_{pi} - \overline{u}_p)}{\sum (w_i - \overline{w})^2}\right)$$

$$\mathbf{plim}\left(\frac{A}{B}\right) = \frac{\mathbf{plim}\,A}{\mathbf{plim}\,B}$$

 $\mathbf{plim}\left(\frac{A}{B}\right) = \frac{\mathbf{plim}\,A}{\mathbf{plim}\,B} \quad \text{if } A \text{ and } B \text{ have probability limits} \\ \text{and plim } B \text{ is not 0.}$ 

We focus on the error term. We would like to use the plim quotient rule. The plim of a quotient is the plim of the numerator divided by the plim of the denominator, provided that both of these limits exist.

$$p = \beta_1 + \beta_2 w + u_{\mu}$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$b_2^{\text{OLS}} = \beta_2 + \frac{\sum (w_i - \overline{w})(u_{pi} - \overline{u}_p)}{\sum (w_i - \overline{w})^2}$$

$$\mathbf{plim}\left(\frac{\sum (w_i - \overline{w})(u_{pi} - \overline{u}_p)}{\sum (w_i - \overline{w})^2}\right)$$

$$\mathbf{plim}\left(\frac{A}{B}\right) = \frac{\mathbf{plim}\,A}{\mathbf{plim}\,B}$$

 $\mathbf{plim}\left(\frac{A}{B}\right) = \frac{\mathbf{plim}\,A}{\mathbf{plim}\,B} \quad \text{if } A \text{ and } B \text{ have probability limits} \\ \text{and plim } B \text{ is not 0.}$ 

However, as the expression stands, the numerator and the denominator of the error term do not have limits. The denominator increases indefinitely as the sample size increases. The nominator has no particular limit.

$$p = \beta_1 + \beta_2 w + u_p$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$b_2^{\text{OLS}} = \beta_2 + \frac{\sum (w_i - \overline{w})(u_{pi} - \overline{u}_p)}{\sum (w_i - \overline{w})^2}$$

$$\operatorname{plim}\left(\frac{\sum (w_i - \overline{w})(u_{pi} - \overline{u}_p)}{\sum (w_i - \overline{w})^2}\right) = \operatorname{plim}\left(\frac{\frac{1}{n}\sum (w_i - \overline{w})(u_{pi} - \overline{u}_p)}{\frac{1}{n}\sum (w_i - \overline{w})^2}\right)$$

$$\operatorname{plim}\left(\frac{A}{B}\right) = \frac{\operatorname{plim} A}{\operatorname{plim} B} \quad \text{if } A$$

 $\operatorname{plim}\left(\frac{A}{B}\right) = \frac{\operatorname{plim} A}{\operatorname{plim} B} \quad \text{if } A \text{ and } B \text{ have probability limits}$  and plim B is not 0.

To deal with this problem, we divide both the numerator and the denominator by n.

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$b_2^{\text{OLS}} = \beta_2 + \frac{\sum (w_i - \overline{w})(u_{pi} - \overline{u}_p)}{\sum (w_i - \overline{w})^2}$$

$$\operatorname{plim}\left(\frac{\sum (w_i - \overline{w})(u_{pi} - \overline{u}_p)}{\sum (w_i - \overline{w})^2}\right) = \operatorname{plim}\left(\frac{\frac{1}{n}\sum (w_i - \overline{w})(u_{pi} - \overline{u}_p)}{\frac{1}{n}\sum (w_i - \overline{w})^2}\right)$$

$$\operatorname{plim} \frac{1}{n} \sum (w_i - \overline{w}) (u_{pi} - \overline{u}_p) = \operatorname{cov}(w, u_p)$$

$$p\lim_{n} \frac{1}{n} \sum_{i} (w_i - \overline{w})^2 = var(w)$$

It can be shown that the limit of the numerator is the covariance of w and  $u_p$  and the limit of the denominator is the variance of w.

structural equations

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$b_2^{\text{OLS}} = \beta_2 + \frac{\sum (w_i - \overline{w})(u_{pi} - \overline{u}_p)}{\sum (w_i - \overline{w})^2}$$

$$\operatorname{plim}\left(\frac{\sum (w_{i} - \overline{w})(u_{pi} - \overline{u}_{p})}{\sum (w_{i} - \overline{w})^{2}}\right) = \operatorname{plim}\left(\frac{\frac{1}{n}\sum (w_{i} - \overline{w})(u_{pi} - \overline{u}_{p})}{\frac{1}{n}\sum (w_{i} - \overline{w})^{2}}\right)$$
$$= \frac{\operatorname{cov}(w, u_{p})}{\operatorname{var}(w)}$$

Hence the numerator and the denominator of the error term have limits and we are entitled to implement the plim quotient rule. We need var(w) to be non-zero, but this will be the case assuming that there is some variation in w.

$$p = \beta_1 + \beta_2 w + u_{\mu}$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\operatorname{plim} b_2^{\operatorname{OLS}} = \beta_2 + \frac{\operatorname{cov}(u_p, w)}{\operatorname{var}(w)}$$

$$\operatorname{cov}(u_{p}, w) = \operatorname{cov}\left(u_{p}, \frac{\alpha_{1} + \alpha_{2}\beta_{1} + \alpha_{3}U + \alpha_{2}u_{p} + u_{w}}{1 - \alpha_{2}\beta_{2}}\right)$$

reduced form equation

$$w = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}$$

We will now derive the limiting value of the numerator. The first step is to substitute for w from its reduced form equation. (*Note*: Here we must use the reduced form equation. If we use the structural equation, we will find ourselves going round in circles.)

structural equations

$$p = \beta_1 + \beta_2 w + u_p$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\operatorname{plim} b_2^{\operatorname{OLS}} = \beta_2 + \frac{\operatorname{cov}(u_p, w)}{\operatorname{var}(w)}$$

$$cov(u_p, w) = cov\left(u_p, \frac{\alpha_1 + \alpha_2\beta_1 + \alpha_3U + \alpha_2u_p + u_w}{1 - \alpha_2\beta_2}\right)$$

$$= \frac{1}{1 - \alpha_2\beta_2} \begin{cases} cov(u_p, [\alpha_1 + \alpha_2\beta_1]) + cov(u_p, \alpha_3U) \\ + cov(u_p, \alpha_2u_p) + cov(u_p, u_w) \end{cases}$$

We use Covariance Rule 1 to decompose the expression.

structural equations

$$p = \beta_1 + \beta_2 w + u_p$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\operatorname{plim} b_2^{\operatorname{OLS}} = \beta_2 + \frac{\operatorname{cov}(u_p, w)}{\operatorname{var}(w)}$$

$$cov(u_p, w) = cov\left(u_p, \frac{\alpha_1 + \alpha_2\beta_1 + \alpha_3U + \alpha_2u_p + u_w}{1 - \alpha_2\beta_2}\right)$$

$$= \frac{1}{1 - \alpha_2\beta_2} \begin{cases} cov(u_p, [\alpha_1 + \alpha_2\beta_1]) + cov(u_p, \alpha_3U) \\ + cov(u_p, \alpha_2u_p) + cov(u_p, u_w) \end{cases}$$

$$= \frac{1}{1 - \alpha_2\beta_2} \left(0 + 0 + \alpha_2 var(u_p) + 0\right) = \frac{\alpha_2\sigma_{u_p}^2}{1 - \alpha_2\beta_2}$$

The first term is 0 because  $(\alpha_1 + \alpha_2 \beta_1)$  is a constant. The second term is 0 because *U* is exogenous and so distributed independently of  $u_p$ 

structural equations

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\operatorname{plim} b_2^{\operatorname{OLS}} = \beta_2 + \frac{\operatorname{cov}(u_p, w)}{\operatorname{var}(w)}$$

$$cov(u_p, w) = cov\left(u_p, \frac{\alpha_1 + \alpha_2\beta_1 + \alpha_3U + \alpha_2u_p + u_w}{1 - \alpha_2\beta_2}\right)$$

$$= \frac{1}{1 - \alpha_2\beta_2} \begin{cases} cov(u_p, [\alpha_1 + \alpha_2\beta_1]) + cov(u_p, \alpha_3U) \\ + cov(u_p, \alpha_2u_p) + cov(u_p, u_w) \end{cases}$$

$$= \frac{1}{1 - \alpha_2\beta_2} \left(0 + 0 + \alpha_2 var(u_p) + 0\right) = \frac{\alpha_2\sigma_{u_p}^2}{1 - \alpha_2\beta_2}$$

The fourth term is 0 if the disturbance terms are distributed independently of each other. This is not necessarily the case but, for simplicity, we will assume it to be true.

structural equations

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\operatorname{plim} b_2^{\operatorname{OLS}} = \beta_2 + \frac{\operatorname{cov}(u_p, w)}{\operatorname{var}(w)}$$

$$cov(u_p, w) = cov\left(u_p, \frac{\alpha_1 + \alpha_2\beta_1 + \alpha_3U + \alpha_2u_p + u_w}{1 - \alpha_2\beta_2}\right)$$

$$= \frac{1}{1 - \alpha_2\beta_2} \begin{cases} cov(u_p, [\alpha_1 + \alpha_2\beta_1]) + cov(u_p, \alpha_3U) \\ + cov(u_p, \alpha_2u_p) + cov(u_p, u_w) \end{cases}$$

$$= \frac{1}{1 - \alpha_2\beta_2} \left(0 + 0 + \alpha_2 var(u_p) + 0\right) = \frac{\alpha_2\sigma_{u_p}^2}{1 - \alpha_2\beta_2}$$

However, the third term is nonzero because the limiting value of a sample variance is the corresponding population variance.

$$p = \beta_1 + \beta_2 w + u_{\mu}$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\operatorname{plim} b_2^{\operatorname{OLS}} = \beta_2 + \frac{\operatorname{cov}(u_p, w)}{\operatorname{var}(w)}$$

$$\operatorname{var}(w) = \operatorname{var}\left(\frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}\right)$$

reduced form equation

$$w = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}$$

If we were interested in obtaining an explicit mathematical expression for the large-sample bias, we would decompose var(w) in the same way, substituting for w from the reduced form, expanding, and then simplifying as best we can.

$$p = \beta_1 + \beta_2 w + u_{\mu}$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p\lim b_2^{OLS} = \beta_2 + \frac{\operatorname{cov}(u_p, w)}{\operatorname{var}(w)}$$

$$\operatorname{cov}(u_p, w) = \frac{\alpha_2 \sigma_{u_p}^2}{1 - \alpha_2 \beta_2}$$

However, usually we are content with determining the sign of the large sample bias, if we can. Since variances are always positive, the sign of the bias will depend on the sign of  $cov(u_p, w)$ .

$$p = \beta_1 + \beta_2 w + u_{\mu}$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\operatorname{plim} b_2^{\operatorname{OLS}} = \beta_2 + \frac{\operatorname{cov}(u_p, w)}{\operatorname{var}(w)}$$

$$\operatorname{cov}(u_p, w) = \frac{\alpha_2 \sigma_{u_p}^2}{1 - \alpha_2 \beta_2}$$

The sign of the bias will depend on the sign of the term  $(1 - \alpha_2 \beta_2)$ , since  $\alpha_2$  must be positive and the variance components are positive.

$$p = \beta_1 + \beta_2 w + u$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = \beta_1 + \beta_2 (\alpha_1 + \alpha_2 p + \alpha_3 U + u_w) + u_p$$
$$(1 - \alpha_2 \beta_2) p = \beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w$$

$$p = \frac{\beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w}{1 - \alpha_2 \beta_2}$$

$$w = \alpha_1 + \alpha_2 (\beta_1 + \beta_2 w + u_p) + \alpha_3 U + u_w$$
$$(1 - \alpha_2 \beta_2) w = \alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w$$

$$w = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}$$

Looking at the reduced form equation for w, w should be a decreasing function of U.  $\alpha_3$ should be negative. So  $(1 - \alpha_2 \beta_2)$  must be positive. We conclude that, in this particular case, the large-sample bias is positive.

$$p = \beta_1 + \beta_2 w + u_p$$

$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\Delta w = \alpha_3 \Delta U$$

In fact,  $(1 - \alpha_2 \beta_2)$  being positive is a condition for the existence of equilibrium in this model. Suppose that the exogenous variable U changed by an amount  $\Delta U$ . The immediate effect on w would be to change it by  $\alpha_3 \Delta U$  (in the opposite direction, since  $\alpha_3 < 0$ ).

$$p = \beta_1 + \beta_2 w + u_p$$

$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\Delta w = \alpha_3 \Delta U$$

$$\Delta p = \beta_2 \Delta w = \beta_2 \alpha_3 \Delta U$$

This would cause p to change by  $\beta_2 \alpha_3 \Delta U$ .

$$p = \beta_1 + \beta_2 w + u_p$$

$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\Delta w = \alpha_3 \Delta U$$

$$\Delta p = \beta_2 \Delta w = \beta_2 \alpha_3 \Delta U$$

$$\Delta w = \alpha_3 \Delta U + \alpha_2 \Delta p$$
$$= (1 + \alpha_2 \beta_2) \alpha_3 \Delta U$$

This would cause a secondary change in w equal to  $\alpha_2\beta_2\alpha_3\Delta U$ .

$$p = \beta_1 + \beta_2 w + u_p$$

$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\Delta w = \alpha_3 \Delta U$$

$$\Delta p = \beta_2 \Delta w = \beta_2 \alpha_3 \Delta U$$

$$\Delta w = \alpha_3 \Delta U + \alpha_2 \Delta p$$
$$= (1 + \alpha_2 \beta_2) \alpha_3 \Delta U$$

$$\Delta p = \beta_2 \Delta w = \beta_2 (1 + \alpha_2 \beta_2) \alpha_3 \Delta U$$

This in turn would cause p to change by a secondary amount  $\alpha_2 \beta_2^2 \alpha_3 \Delta U$ .

$$p = \beta_1 + \beta_2 w + u_p$$

$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\Delta w = \alpha_3 \Delta U$$

$$\Delta p = \beta_2 \Delta w = \beta_2 \alpha_3 \Delta U$$

$$\Delta w = \alpha_3 \Delta U + \alpha_2 \Delta p$$
$$= (1 + \alpha_2 \beta_2) \alpha_3 \Delta U$$

$$\Delta p = \beta_2 \Delta w = \beta_2 (1 + \alpha_2 \beta_2) \alpha_3 \Delta U$$

$$\Delta w = \alpha_3 \Delta U + \alpha_2 \Delta p$$
$$= \left(1 + \alpha_2 \beta_2 + \alpha_2^2 \beta_2^2\right) \alpha_3 \Delta U$$

This would cause w to change by a further amount  $\alpha_2^2 \beta_2^2 \alpha_3 \Delta U$ .

$$p = \beta_1 + \beta_2 w + u_p$$

$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\Delta w = \alpha_3 \Delta U$$

$$\Delta p = \beta_2 \Delta w = \beta_2 \alpha_3 \Delta U$$

$$\Delta w = \alpha_3 \Delta U + \alpha_2 \Delta p$$
$$= (1 + \alpha_2 \beta_2) \alpha_3 \Delta U$$

$$\Delta p = \beta_2 \Delta w = \beta_2 (1 + \alpha_2 \beta_2) \alpha_3 \Delta U$$

$$\Delta w = \alpha_3 \Delta U + \alpha_2 \Delta p$$

$$= \left(1 + \alpha_2 \beta_2 + \alpha_2^2 \beta_2^2\right) \alpha_3 \Delta U$$

$$= \left(1 + \alpha_2 \beta_2 + \alpha_2^2 \beta_2^2 + \alpha_2^3 \beta_2^3 + ...\right) \alpha_3 \Delta U$$

And so on and so forth. The total change will be finite only if  $\alpha_2\beta_2 < 1$ . Otherwise the process would be explosive, which is implausible.

$$p = \beta_1 + \beta_2 w + u_1$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p\lim b_2^{OLS} = \beta_2 + \frac{\operatorname{cov}(u_p, w)}{\operatorname{var}(w)}$$

$$\operatorname{cov}(u_p, w) = \frac{\alpha_2 \sigma_{u_p}^2}{1 - \alpha_2 \beta_2}$$

Either way, we have demonstrated that  $1 - \alpha_2 \beta_2 > 0$  and hence that, in this case, the bias is positive. Note that one cannot generalize about the direction of simultaneous equations bias. It depends on the structure of the model.

structural equations

$$p = \beta_1 + \beta_2 w + u_p$$

$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

In the previous slideshow we determined analytically the large-sample simultaneous equations bias for the price inflation / wage inflation model. Next we will look at the bias graphically.

$$p = \beta_1 + \beta_2 w + u_{\mu}$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\operatorname{var}(w) = \operatorname{var}\left(\frac{\alpha_1 + \alpha_2 \beta_1}{1 - \alpha_2 \beta_2} + \frac{\alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}\right) = \operatorname{var}\left(\frac{\alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}\right)$$

$$w = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}$$

For this, it will be necessary to obtain an explicit expression for var(w). Variances are unaffected by additive constants, so the first part of the expression may be dropped.

structural equations

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\operatorname{var}(w) = \operatorname{var}\left(\frac{\alpha_{1} + \alpha_{2}\beta_{1}}{1 - \alpha_{2}\beta_{2}} + \frac{\alpha_{3}U + \alpha_{2}u_{p} + u_{w}}{1 - \alpha_{2}\beta_{2}}\right) = \operatorname{var}\left(\frac{\alpha_{3}U + \alpha_{2}u_{p} + u_{w}}{1 - \alpha_{2}\beta_{2}}\right)$$

$$= \left\{\frac{1}{(1 - \alpha_{2}\beta_{2})^{2}}\right\} \left\{ \operatorname{var}(\alpha_{3}U) + \operatorname{var}(\alpha_{2}u_{p}) + \operatorname{var}(u_{w}) + 2\operatorname{cov}(\alpha_{3}U, \alpha_{2}u_{p}) + 2\operatorname{cov}(\alpha_{3}U, \alpha_{2}u_{p}) + 2\operatorname{cov}(\alpha_{2}u_{p}, u_{w}) + 2\operatorname{cov}(\alpha_{2}$$

Using Variance Rule 1, the numerator decomposes into three variances and three covariances. The denominator is a constant common factor and may be taken, squared, outside the expression using Variance Rule 2.

structural equations

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$var(w) = var\left(\frac{\alpha_{1} + \alpha_{2}\beta_{1}}{1 - \alpha_{2}\beta_{2}} + \frac{\alpha_{3}U + \alpha_{2}u_{p} + u_{w}}{1 - \alpha_{2}\beta_{2}}\right) = var\left(\frac{\alpha_{3}U + \alpha_{2}u_{p} + u_{w}}{1 - \alpha_{2}\beta_{2}}\right)$$

$$= \left\{\frac{1}{(1 - \alpha_{2}\beta_{2})^{2}}\right\} \begin{cases} var(\alpha_{3}U) + var(\alpha_{2}u_{p}) + var(u_{w}) \\ + 2cov(\alpha_{3}U, u_{w}) + 2cov(\alpha_{3}U, \alpha_{2}u_{p}) \\ + 2cov(\alpha_{2}u_{p}, u_{w}) \end{cases}$$

$$= \frac{\alpha_{3}^{2}\sigma_{U}^{2} + \alpha_{2}^{2}\sigma_{u_{p}}^{2} + \sigma_{u_{w}}^{2}}{(1 - \alpha_{2}\beta_{2})^{2}}$$

The covariances are all 0 on the assumption that U,  $u_p$ , and  $u_w$  are distributed independently of each other. Thus the numerator consists of the variance expressions. Remember that we have to square  $\alpha_3$  and  $\alpha_2$  when we take them out of the variance expressions.

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\operatorname{plim} b_2^{\operatorname{OLS}} = \beta_2 + \frac{\operatorname{cov}(u_p, w)}{\operatorname{var}(w)}$$

$$\operatorname{cov}(u_p, w) = \frac{\alpha_2 \sigma_{u_p}^2}{1 - \alpha_2 \beta_2}$$

$$var(w) = \frac{\alpha_3^2 \sigma_U^2 + \alpha_2^2 \sigma_{u_p}^2 + \sigma_{u_w}^2}{(1 - \alpha_2 \beta_2)^2}$$

plim 
$$b_2^{\text{OLS}} = \beta_2 + (1 - \alpha_2 \beta_2) \frac{\alpha_2 \sigma_{u_p}^2}{\alpha_3^2 \sigma_U^2 + \alpha_2^2 \sigma_{u_p}^2 + \sigma_{u_w}^2}$$

Hence we obtain an explicit expression for the plim of the slope coefficient.

structural equations

$$p = \beta_1 + \beta_2 w + u_{\mu}$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\begin{aligned}
\mathbf{plim} \, b_2^{\text{OLS}} &= \beta_2 + \left(1 - \alpha_2 \beta_2\right) \frac{\alpha_2 \sigma_{u_p}^2}{\alpha_3^2 \sigma_U^2 + \alpha_2^2 \sigma_{u_p}^2 + \sigma_{u_w}^2} \\
&= \beta_2 + \left(\frac{1}{\alpha_2} - \beta_2\right) \frac{\alpha_2^2 \sigma_{u_p}^2}{\alpha_3^2 \sigma_U^2 + \alpha_2^2 \sigma_{u_p}^2 + \sigma_{u_w}^2}
\end{aligned}$$

To look at the bias graphically, it is helpful to rearrange the expression.

structural equations

$$p = \beta_1 + \beta_2 w + u_{\mu}$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\begin{aligned}
\mathbf{plim} \, b_{2}^{\text{OLS}} &= \beta_{2} + \left(1 - \alpha_{2} \beta_{2}\right) \frac{\alpha_{2} \sigma_{u_{p}}^{2}}{\alpha_{3}^{2} \sigma_{U}^{2} + \alpha_{2}^{2} \sigma_{u_{p}}^{2} + \sigma_{u_{w}}^{2}} \\
&= \beta_{2} + \left(\frac{1}{\alpha_{2}} - \beta_{2}\right) \frac{\alpha_{2}^{2} \sigma_{u_{p}}^{2}}{\alpha_{3}^{2} \sigma_{U}^{2} + \alpha_{2}^{2} \sigma_{u_{p}}^{2} + \sigma_{u_{w}}^{2}} \\
&= \beta_{2} \left(1 - \frac{\alpha_{2}^{2} \sigma_{u_{p}}^{2}}{\alpha_{3}^{2} \sigma_{U}^{2} + \alpha_{2}^{2} \sigma_{u_{p}}^{2} + \sigma_{u_{w}}^{2}}\right) + \frac{1}{\alpha_{2}} \left(\frac{\alpha_{2}^{2} \sigma_{u_{p}}^{2}}{\alpha_{3}^{2} \sigma_{U}^{2} + \alpha_{2}^{2} \sigma_{u_{p}}^{2} + \sigma_{u_{w}}^{2}}\right)
\end{aligned}$$

We have now gathered the terms involving  $\beta_2$  together.

structural equations

$$p = \beta_1 + \beta_2 w + u_p$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\begin{aligned} \text{plim } b_2^{\text{OLS}} &= \beta_2 + \left(1 - \alpha_2 \beta_2\right) \frac{\alpha_2 \sigma_{u_p}^2}{\alpha_3^2 \sigma_U^2 + \alpha_2^2 \sigma_{u_p}^2 + \sigma_{u_w}^2} \\ &= \beta_2 + \left(\frac{1}{\alpha_2} - \beta_2\right) \frac{\alpha_2^2 \sigma_{u_p}^2}{\alpha_3^2 \sigma_U^2 + \alpha_2^2 \sigma_{u_p}^2 + \sigma_{u_w}^2} \\ &= \beta_2 \left(1 - \frac{\alpha_2^2 \sigma_{u_p}^2}{\alpha_3^2 \sigma_U^2 + \alpha_2^2 \sigma_{u_p}^2 + \sigma_{u_w}^2}\right) + \frac{1}{\alpha_2} \left(\frac{\alpha_2^2 \sigma_{u_p}^2}{\alpha_3^2 \sigma_U^2 + \alpha_2^2 \sigma_{u_p}^2 + \sigma_{u_w}^2}\right) \\ &= \beta_2 \left(\frac{\alpha_3^2 \sigma_U^2 + \sigma_{u_w}^2}{\alpha_3^2 \sigma_U^2 + \alpha_2^2 \sigma_{u_p}^2 + \sigma_{u_w}^2}\right) + \frac{1}{\alpha_2} \left(\frac{\alpha_2^2 \sigma_{u_p}^2}{\alpha_3^2 \sigma_U^2 + \alpha_2^2 \sigma_{u_p}^2 + \sigma_{u_w}^2}\right) \end{aligned}$$

Thus we see that the limiting value of the OLS estimator is a weighted average of the true value  $\beta_2$  and  $1/\alpha_2$ .

$$p = \beta_1 + \frac{\beta_2}{\beta_2} w + u_1$$

$$p = \beta_1 + \frac{\beta_2}{\beta_2}w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\mathbf{plim} \, b_{2}^{\text{OLS}} = \frac{\beta_{2}}{\alpha_{3}^{2} \sigma_{U}^{2} + \alpha_{2}^{2} \sigma_{u_{p}}^{2} + \sigma_{u_{w}}^{2}} + \frac{1}{\alpha_{2}} \left( \frac{\alpha_{2}^{2} \sigma_{u_{p}}^{2}}{\alpha_{3}^{2} \sigma_{U}^{2} + \alpha_{2}^{2} \sigma_{u_{p}}^{2} + \sigma_{u_{w}}^{2}} \right) + \frac{1}{\alpha_{2}} \left( \frac{\alpha_{2}^{2} \sigma_{u_{p}}^{2}}{\alpha_{3}^{2} \sigma_{U}^{2} + \alpha_{2}^{2} \sigma_{u_{p}}^{2} + \sigma_{u_{w}}^{2}} \right)$$

$$p = \frac{\beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w}{1 - \alpha_2 \beta_2}$$

$$w = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}$$

Variations in *U* cause variations in *p* and *w*, the change in *p* being  $\beta_2$  times the change in *w*. Such movements trace out the true relationship.

$$p = \beta_1 + \beta_2 w + u_1$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$\mathbf{plim}\,b_{2}^{\text{OLS}} = \beta_{2} \left( \frac{\alpha_{3}^{2} \sigma_{U}^{2} + \sigma_{u_{w}}^{2}}{\alpha_{3}^{2} \sigma_{U}^{2} + \alpha_{2}^{2} \sigma_{u_{p}}^{2} + \sigma_{u_{w}}^{2}} \right) + \frac{1}{\alpha_{2}} \left( \frac{\alpha_{2}^{2} \sigma_{u_{p}}^{2}}{\alpha_{3}^{2} \sigma_{U}^{2} + \alpha_{2}^{2} \sigma_{u_{p}}^{2} + \sigma_{u_{w}}^{2}} \right)$$

$$p = \frac{\beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w}{1 - \alpha_2 \beta_2}$$

$$w = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}$$

Likewise, variations in  $u_p$  cause variations in p and w, the change in p being  $1/\alpha_2$  times the change in w.

structural equations

$$p = \beta_1 + \beta_2 w + u_p$$

$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = 1.5 + 0.5w + u_p$$

$$w = 2.5 + 0.5p - 0.4U$$

n = 20.

U = 2 to 6.75 in steps of 0.25. Variance 2.08.

 $u_p \sim \text{iid } N(0, 0.64).$ 

 $u_w$  dropped.

We will illustrate this with a Monte Carlo model, choosing parameter values as shown above. The sample size 1s 20 and the values of U are 2, to 6.75 in steps of 0.25.

$$p = \beta_1 + \beta_2 w + u_p$$

$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = 1.5 + 0.5w + u_p$$

$$w = 2.5 + 0.5p - 0.4U$$

$$n = 20.$$

U = 2 to 6.75 in steps of 0.25. Variance 2.08.

 $u_p \sim \text{iid } N(0, 0.64).$ 

 $u_w$  dropped.

reduced form equation

$$p = \frac{\beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + u_p + \beta_2 u_w}{1 - \alpha_2 \beta_2} = 3.67 - 0.27U + 1.33u_p$$

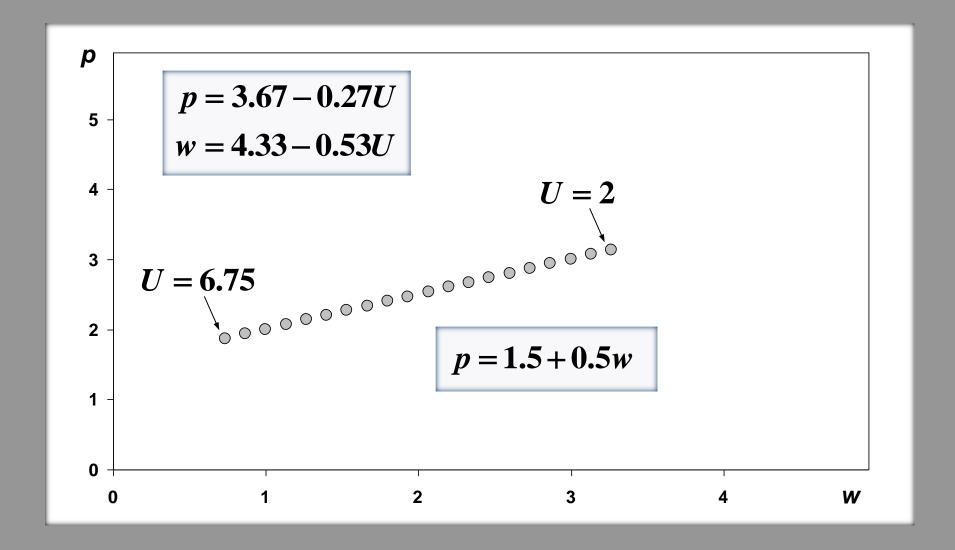
$$= 3.67 - 0.27U + 1.33u_{\mu}$$

reduced form equation

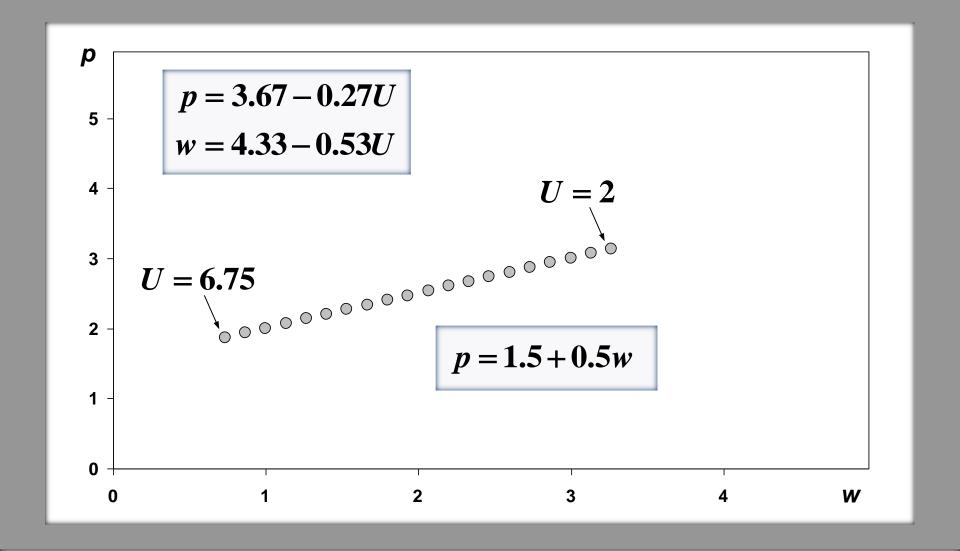
$$w = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}$$

$$= 4.33 - 0.53U + 0.67u_p$$

Given the values of the parameters, the reduced form equations are as shown.



This is what we would see if there were no disturbance terms in the model.



Note that the values of p and w in each observation are jointly determined by the value of U in that observation. A change in U changes p by only half the amount it changes w. Thus the slope is 0.5.

$$p = \beta_1 + \beta_2 w + u_{\mu}$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = 1.5 + 0.5w + u_{p}$$

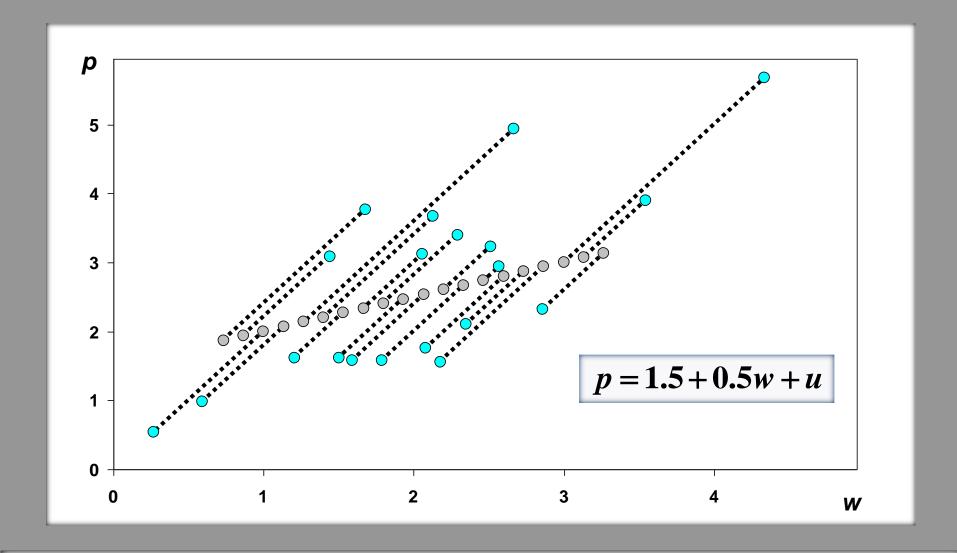
$$p = 1.5 + 0.5w + u_n$$
  $w = 2.5 + 0.5p - 0.4U$ 

$$\mathbf{plim}\,b_{2}^{\text{OLS}} = \beta_{2} \left( \frac{\alpha_{3}^{2}\sigma_{U}^{2} + \sigma_{u_{w}}^{2}}{\alpha_{3}^{2}\sigma_{U}^{2} + \alpha_{2}^{2}\sigma_{u_{p}}^{2} + \sigma_{u_{w}}^{2}} \right) + \frac{1}{\alpha_{2}} \left( \frac{\alpha_{2}^{2}\sigma_{u_{p}}^{2}}{\alpha_{3}^{2}\sigma_{U}^{2} + \alpha_{2}^{2}\sigma_{u_{p}}^{2} + \sigma_{u_{w}}^{2}} \right)$$

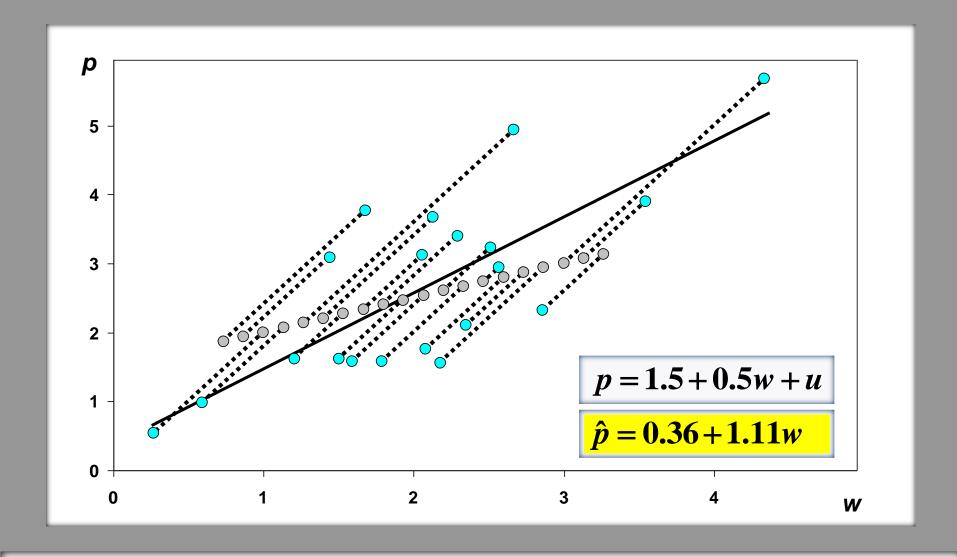
$$p = \frac{\beta_1 + \alpha_1 \beta_2 + \alpha_3 \beta_2 U + \mathbf{u}_p + \beta_2 \mathbf{u}_w}{1 - \alpha_2 \beta_2}$$

$$w = \frac{\alpha_1 + \alpha_2 \beta_1 + \alpha_3 U + \alpha_2 u_p + u_w}{1 - \alpha_2 \beta_2}$$

At the same time, p and w are both affected by variations in  $u_p$ , the change in p being  $1/\alpha_2$ times the change in w.



As a consequence the actual observations are shifted away from the true relationship up or down along lines with slope  $1/\alpha_2$ , the dotted lines in the graph. Since  $\alpha_2$  is 0.5,  $1/\alpha_2$  is 2.0.



The regression line thus overestimates  $\beta_2$ .

structural equations

$$p = \beta_1 + \beta_2 w + u_{\mu}$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = 1.5 + 0.5w + u_p$$

$$p = 1.5 + 0.5w + u_p$$
  $w = 2.5 + 0.5p - 0.4U$ 

$$\begin{aligned}
\text{plim } b_2^{\text{OLS}} &= \beta_2 \left( \frac{\alpha_3^2 \sigma_U^2 + \sigma_{u_w}^2}{\alpha_3^2 \sigma_U^2 + \alpha_2^2 \sigma_{u_p}^2 + \sigma_{u_w}^2} \right) + \frac{1}{\alpha_2} \left( \frac{\alpha_2^2 \sigma_{u_p}^2}{\alpha_3^2 \sigma_U^2 + \alpha_2^2 \sigma_{u_p}^2 + \sigma_{u_w}^2} \right) \\
&= 0.5 \left( \frac{0.16 \times 2.08}{0.16 \times 2.08 + 0.25 \times 0.64} \right) \\
&+ 2.0 \left( \frac{0.25 \times 0.64}{0.16 \times 2.08 + 0.25 \times 0.64} \right) & \sigma_U^2 &= 2.08 \\
\sigma_{u_p}^2 &= 0.64 \\
\sigma_{u_w}^2 &= 0
\end{aligned}$$

We will calculate the large-sample bias in the slope coefficient. U and  $u_p$  were chosen so that they had population variances 2.08 and 0.64, respectively.

structural equations

$$p = \beta_1 + \beta_2 w + u_1$$

$$p = \beta_1 + \beta_2 w + u_p \qquad w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = 1.5 + 0.5w + u_p$$

$$p = 1.5 + 0.5w + u_p$$
  $w = 2.5 + 0.5p - 0.4U$ 

$$\begin{aligned}
\text{plim } b_2^{\text{OLS}} &= \beta_2 \left( \frac{\alpha_3^2 \sigma_U^2 + \sigma_{u_w}^2}{\alpha_3^2 \sigma_U^2 + \alpha_2^2 \sigma_{u_p}^2 + \sigma_{u_w}^2} \right) + \frac{1}{\alpha_2} \left( \frac{\alpha_2^2 \sigma_{u_p}^2}{\alpha_3^2 \sigma_U^2 + \alpha_2^2 \sigma_{u_p}^2 + \sigma_{u_w}^2} \right) \\
&= 0.5 \left( \frac{0.16 \times 2.08}{0.16 \times 2.08 + 0.25 \times 0.64} \right) \\
&+ 2.0 \left( \frac{0.25 \times 0.64}{0.16 \times 2.08 + 0.25 \times 0.64} \right) & \sigma_U^2 &= 2.08 \\
&= 0.99 & \sigma_{u_p}^2 &= 0.64 \\
&= 0.99
\end{aligned}$$

The large sample bias is 0.99. Our regression estimate, 1.11, was a little higher.

structural
equations

$$p = \beta_1 + \beta_2 w + u_p$$

$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = 1.5 + \frac{0.5}{0.5}w + u_p$$

$$w = 2.5 + 0.5p - 0.4U$$

	<b>b</b> <sub>1</sub>	s.e.( <i>b</i> <sub>1</sub> )	$b_2$	s.e.( <i>b</i> <sub>2</sub> )	
1	0.36	0.49	1.11	0.22	
2	0.45	0.38	1.06	0.17	
3	0.65	0.27	0.94	0.12	
4	0.41	0.39	0.98	0.19	
5	0.46	0.92	0.77	0.22	
6	0.26	0.35	1.09	0.16	
7	0.31	0.39	1.00	0.19	
8	1.06	0.38	0.82	0.16	
9	-0.08	0.36	1.16	0.18	
10	1.12	0.43	0.69	0.20	

Here are the results for 10 samples in this simulation. The slope coefficient was overestimated every time and does appear to be distributed around its plim, 0.99.

structural
equations

$$p = \beta_1 + \beta_2 w + u_p$$

$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = 1.5 + 0.5w + u_p$$

$$w = 2.5 + 0.5p - 0.4U$$

	<b>b</b> <sub>1</sub>	s.e.( <i>b</i> <sub>1</sub> )	$b_2$	s.e.( <i>b</i> <sub>2</sub> )	
1	0.36	0.49	1.11	0.22	
2	0.45	0.38	1.06	0.17	
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7	0.31	0.39	1.00	0.19	
8	1.06	0.38	0.82	0.16	
9	<b>-0.08</b>	0.36	1.16	0.18	
10	1.12	0.43	0.69	0.20	

Because the slope coefficient was overestimated, the intercept was underestimated every time.

structural
equations

$$p = \beta_1 + \beta_2 w + u_p$$

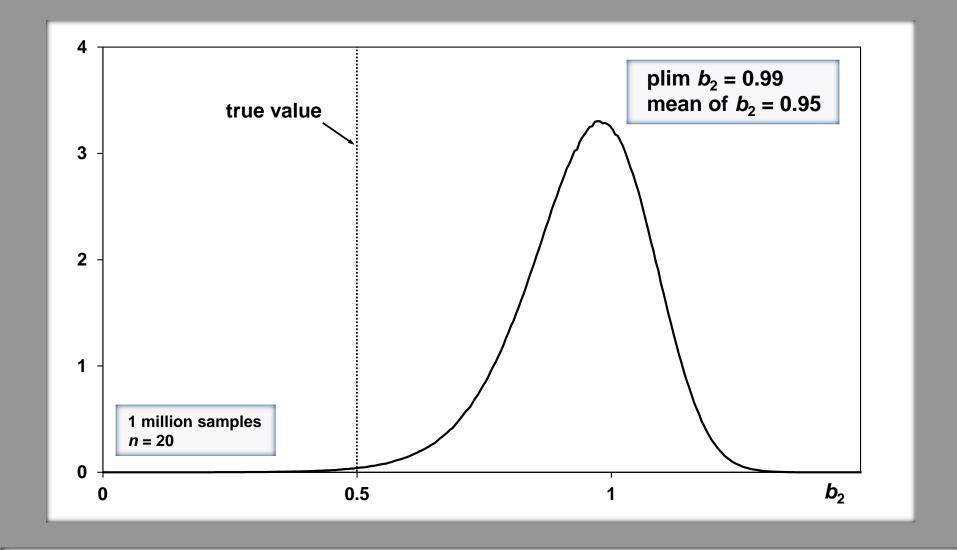
$$w = \alpha_1 + \alpha_2 p + \alpha_3 U + u_w$$

$$p = 1.5 + 0.5w + u_p$$

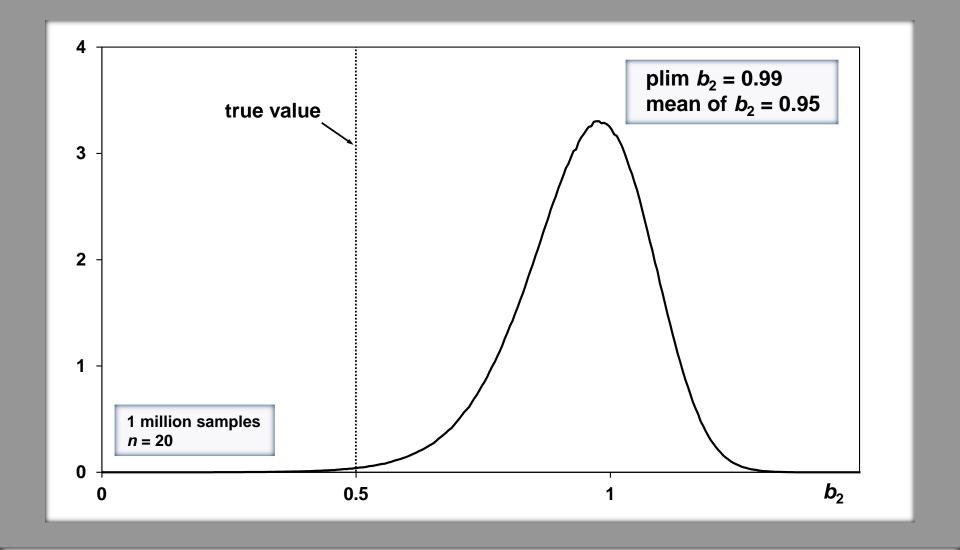
$$w = 2.5 + 0.5p - 0.4U$$

	<b>b</b> <sub>1</sub>	s.e.( <i>b</i> <sub>1</sub> )	<b>b</b> <sub>2</sub>	s.e.( <i>b</i> <sub>2</sub> )
1	0.36	0.49	1.11	0.22
2	0.45	0.38	1.06	0.17
3	0.65	0.27	0.94	0.12
4	0.41	0.39	0.98	0.19
5	0.46	0.92	0.77	0.22
6	0.26	0.35	1.09	0.16
7	0.31	0.39	1.00	0.19
8	1.06	0.38	0.82	0.16
9	-0.08	0.36	1.16	0.18
10	1.12	0.43	0.69	0.20

No attention should be paid to the standard errors because they are invalidated by the simultaneous equations bias.



The chart plots the distribution of the slope coefficient for 1 million samples. Almost all the estimates are above the true value of 0.5, confirming the large-sample analysis.



The plim (0.99) in this case provides a good guide to the size of the bias since the mean of the distribution is 0.95, fairly close to the plim even though, with only 20 observations, each sample was quite small.

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# Q&A





# Thank you very much! Vielen Dank!





